

# Efficient and Optimally Secure Key-Length Extension for Block Ciphers via Randomized Cascading\*

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**Abstract.** We consider the question of efficiently extending the key length of block ciphers. To date, the approach providing highest security is triple encryption (used e.g. in Triple-DES), which was proved to have roughly  $\kappa + \min\{n/2, \kappa/2\}$  bits of security when instantiated with ideal block ciphers with key length  $\kappa$  and block length  $n$ , at the cost of three block-cipher calls per message block.

This paper presents a new practical key-length extension scheme exhibiting  $\kappa + n/2$  bits of security – hence improving upon the security of triple encryption – solely at the cost of *two* block cipher calls and a key of length  $\kappa + n$ . We also provide matching generic attacks showing the optimality of the security level achieved by our approach with respect to a general class of two-query constructions.

**Keywords.** Block ciphers, Cascade encryption, Provable security.

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# 1 Introduction

## 1.1 Key-Length Extension for Block Ciphers

Several practical block cipher designs have been proposed over the last decades and have been the object of extensive cryptanalytic efforts. Examples include DES [1], IDEA [19], BLOWFISH [28], and the currently in-use AES [4]. Within applications, we typically demand that these block ciphers are a good *pseudorandom permutation* (PRP), i.e., in the eyes of a computationally bounded attacker, they behave as a randomly chosen permutation under a random secret key. For instance, PRP security of the underlying block cipher is necessary to infer security of all modes of operations for message encryption (such as counter-mode and CBC encryption [8]) as well as of message authentication codes like CBC-MAC [9] and PMAC [12].

In practice, we define the PRP security level of a block cipher as the complexity required to distinguish it from a random permutation with non-negligible advantage. The *key length*  $\kappa$  of a block cipher crucially limits the achievable security level, since the secret key  $K$  can be recovered given black-box access to  $E(K, \cdot)$  evaluating  $E(\cdot, \cdot)$  approximately  $2^\kappa$  times; obviously, this also yields a PRP distinguishing adversary with equal complexity. Such weakness is *generic*, in the sense that it only depends on  $\kappa$ , and even an *ideal block cipher* suffers from the same attack.<sup>5</sup> In contrast, no real dependency exists between security and the block length  $n$  of a block cipher: No generic attack faster than  $2^\kappa$  exists even if  $n = 1$ . In the following, let us refer to a block cipher with key and block lengths  $\kappa$  and  $n$ , respectively, as a  $(\kappa, n)$ -block cipher.

**KEY LENGTH EXTENSION.** With a continuous increase of the availability of computing resources, the role of the key length has hence never been more important. Key lengths of say fewer than 64 bits are no longer sufficient to ensure security, making key recovery a matter of a few hours even on modest architectures. This is a serious problem for legacy designs such as DES which have very short keys of length 56 bits, but which otherwise do not seem to present significant non-generic security weaknesses. Constructions based on DES also remain very attractive because of its short block length  $n = 64$  which allows enciphering short inputs. This is for example crucial in current applications in the financial industry, such as the EMV standard [6], where the block cipher is applied to PIN numbers, which are very short.

The above described situation motivates the problem of *key-length extension*, which is the main object of this paper: We seek for very efficient constructions provably transforming any  $(\kappa, n)$ -block cipher  $E$  into a  $(\kappa', n)$ -block cipher  $E'$  with both  $\kappa' > \kappa$  and higher PRP security, i.e., the PRP security of  $E'$  should be higher than  $2^\kappa$  whenever  $E$  does not exhibit any non-generic weaknesses. We aim both at providing very efficient approaches to key length extension and at understanding the optimal security achievable by such constructions. Our main contribution will be a new and very efficient two-call key-length extension method outperforming the efficiency of existing solutions by a large margin, and achieving security levels which we prove optimal, and which are comparable (and even better) than those of earlier, less efficient, designs.

**IDEAL CIPHER MODEL.** In our proofs, we model the absence of generic weaknesses of the underlying block cipher by analyzing our constructions when instantiated with an ideal block cipher  $\mathbf{E}$ . In this model, complexity is measured in terms of the number of queries to  $\mathbf{E}$  (so-called *ideal block cipher queries*) and to either  $E'$  under a random  $\kappa'$ -bit key or the given random permutation (we refer to these as *construction queries*). It should be noted that proving security of key-length extension in the ideal cipher model implies absence of generic attacks, treating the underlying cipher as a

<sup>5</sup> As usual, an *ideal block cipher*  $\mathbf{E} : \{0, 1\}^\kappa \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  is the system associating with each key  $k \in \{0, 1\}^\kappa$  an independent randomly chosen permutation  $E(k, \cdot)$  and allowing the adversary to learn  $E(k, x)$  and  $E^{-1}(k, y)$  for  $k, x, y$  of her choice.

black-box, and as we will explain in the next section, all attacks on existing schemes are indeed generic.

## 1.2 Existing Approaches to Key-Length Extension

The short key length  $\kappa = 56$  of DES has constituted the main motivation behind previous work on key-length extension. However, we stress that all previous constructions are generic, and can be applied to *any* block cipher with short keys, hence extending the applicability of these results (as well as the results of this paper) way beyond the specific case of DES.

A first proposal called DESX (due to Rivest) stretches the key length of DES employing a technique called *key whitening* (this approach was later used by Even and Mansour [15]): It is defined such that

$$\text{DESX}_{k_i, k_o, k}(m) = k_o \oplus \text{DES}_k(k_i \oplus m)$$

for all  $m, k_i, k_o \in \{0, 1\}^{64}$  and  $k \in \{0, 1\}^{56}$ . DESX can be generalized to a generic transformation from a  $(\kappa, n)$ -block cipher to a  $(\kappa + 2n, n)$ -block cipher whose security was studied by Kilian and Rogaway [18]: They proved that any successful PRP distinguishing attack requires  $2^{\frac{\kappa+n}{2}}$  queries.<sup>6</sup> They also observe that the same key may be used in both whitening steps (i.e.,  $k_i = k_o$ ) and provide an attack using  $2^{\max\{\kappa, n\}}$  queries.

An alternative to whitening is *cascading* (or cascade encryption), i.e., sequentially composing  $\ell$  block-cipher calls with usually different keys. (This is referred to as a cascade of length  $\ell$ .) It is well known that a cascade of length two does not substantially increase security due to the meet-in-the-middle attack [13]. (Even though a security increase in terms of distinguishing advantage is achieved for low attack complexities [7].) The security properties of a cascade of different ciphers was studied by Even and Goldreich [14] showing that a cascade is at least as strong as the strongest of the ciphers used; and by Maurer and Massey [23] proving that it is at least as secure as the *first* cipher of the cascade, however in a more general attack model.

The meet-in-the-middle attack makes triple encryption the shortest cascade with a potential for significant security gain and indeed it has found widespread usage as *Triple-DES* (3DES) [2,3,5], where given keys  $k_1, k_2, k_3 \in \{0, 1\}^{56}$ , a 64-bit message  $m$  is mapped to

$$3\text{DES}_{k_1, k_2, k_3}(m) = \text{DES}_{k_1}(\text{DES}_{k_2}(\text{DES}_{k_3}(m))) .$$

(A variant with shorter keys  $3\text{DES}'_{k_1, k_2}(m) = \text{DES}_{k_1}(\text{DES}_{k_2}^{-1}(\text{DES}_{k_1}(m)))$  is also sometimes used.) For 3DES (and a variant of  $3\text{DES}'$  with independent keys), Bellare and Rogaway [11] and subsequently Gaži and Maurer [16] have shown security up to roughly  $2^{\kappa + \min\{n, \kappa\}/2}$  queries when DES is replaced by an ideal block cipher. For the case of DES parameters, their result gives concretely security up to  $2^{78}$  queries, whereas the best known attack due to Lucks [21] shows that no security better than  $2^{90}$  can be expected. (It should also be noted that the proof of [16] extends to prove that longer cascades can achieve better security for short keys.)

We emphasize that despite the availability of modern designs with built-in larger keys (e.g.,  $\kappa \in \{128, 192, 256\}$  for AES), Triple-DES remains nowadays popular, not only because of backwards compatibility, but also because its short block size ( $n = 64$  vs.  $n \geq 128$  for AES) is well suited to applications enciphering short inputs such as personal-identification numbers (PINs). For example, it is the basis of the EMV standard for PIN-based authentication of debit and credit card transactions [6]. However, the use of three calls per processed message block is widely considered a drawback within applications which we address and solve in this paper.

<sup>6</sup> Their result is in fact more fine-grained, as they show that  $2^\rho$  construction and  $2^{\kappa+n-\rho}$  ideal block cipher queries, respectively, are necessary for all integers  $\rho$ ; while different bounds for both query types are sometimes justified, we adopt a (by now more standard) worst-case approach only bounding the *sum* of both query numbers.

OTHER RELATED WORK. It is worth mentioning that several works have studied cascading-based security amplification of block ciphers only assumed to satisfy weaker forms of PRP security, both in the information-theoretic [32,24,25,17] as well as in the computational settings [20,26,31]. These results however consider an orthogonal model to ours and are hence incomparable.

### 1.3 Our Results

None of the above efficient constructions provably achieves security beyond  $2^{\kappa+\min\{\kappa,n\}/2}$ , and such security is achieved only at the fairly expensive cost of at least three block-cipher calls per message block. This paper aims at improving the efficiency-security trade-off in key-length extension. We ask the following question: Suppose that we only consider constructions making at most *two* calls to the underlying cipher. *What is the best security level we are expected to achieve?*

BETTER SECURITY AND BETTER EFFICIENCY. Quite surprisingly, our main result (presented in Section 4) exposes a “win-win” situation: We devise a *two-call* construction of a  $(\kappa + n, n)$ -block cipher from any  $(\kappa, n)$ -block cipher with security  $2^{\kappa+n/2}$  in the ideal block cipher model, i.e., the obtained security is *higher* than that of existing three-call designs studied in [11,16].<sup>7</sup> Our construction – which we refer to as the *double XOR-cascade* (2XOR) – is obtained by careful insertion of two randomization steps (with the *same* key value) to a related-key version of double encryption. Concretely, we map each  $n$ -bit message  $m$  to

$$2\text{XOR}_{k,z}(m) = E_{\tilde{k}}(E_k(m \oplus z) \oplus z)$$

for all key values  $k \in \{0,1\}^\kappa$  and  $z \in \{0,1\}^n$ , and where  $\tilde{k}$  is, for example, obtained from  $k$  by flipping one individual bit.

We note that the key length is comparable to the one of the two-key variant of 3DES (assuming  $\kappa \approx n$ ). Intuitively, our construction requires some mild form of related-key security [10] which we obtain for free when the underlying block cipher is ideal, but may be a concern in practice. However, it should be noted that an alternative version of the construction where  $\tilde{k}$  is replaced by an independent and unrelated key value  $k'$  achieves the same security level at the cost of a longer  $(2\kappa + n)$ -bit key, which is for instance still shorter than in DESX with independent whitening keys (for DES parameters).

The core of our security proof (cf. Theorem 3) is a technical argument of independent interest: Namely, we prove that it is hard to distinguish two random, independent, permutations  $\pi_1, \pi_2$  on the  $n$ -bit strings from two randomly chosen permutations  $\pi_1, \pi_2$  with the property that  $\pi_2(\pi_1(x \oplus Z) \oplus Z) = x$  for all  $x$  and a random secret value  $Z$  even if we are allowed arbitrary queries to each of  $\pi_1, \pi_2, \pi_1^{-1}$ , and  $\pi_2^{-1}$ . This fact yields our main theorem by a careful adaptation of the techniques from [11,16] to take into account both randomization and the use of related keys.

GENERIC ATTACKS AND OPTIMALITY. With the above result at hand, it is legitimate to ask whether we should expect two-call constructions with even better security: In Section 3, we answer this in the negative, at least for a class of natural constructions.

As a warm up of independent interest, we confirm that only much weaker security can be achieved by a one-call construction: Regardless of the amount of key material employed in the construction, an attack with query complexity  $2^{\max\{\kappa,n\}}$  always exists (using memory  $2^{\max\{\kappa,n\}}$ ),<sup>8</sup> showing the optimality of DESX-like constructions in the case  $\kappa = n$ . We then turn to two-call

<sup>7</sup> In fact, our construction tolerates arbitrarily many construction queries (i.e., up to  $2^n$ ) and  $2^{\kappa+n/2}$  ideal block cipher queries. However, we stress that in all practically relevant cases  $\kappa \geq n/2$ , hence we tacitly assume this property throughout the paper.

<sup>8</sup> More precisely, our attack requires roughly  $2^\kappa$  ideal block cipher queries and  $2^n$  construction queries.

| construction                     | # of queries | key length   | log of the number of queries   |                         |
|----------------------------------|--------------|--------------|--------------------------------|-------------------------|
|                                  |              |              | security lower bound           | best known attack       |
| $(\kappa, n)$ -block cipher      | 1            | $\kappa$     | $\kappa$                       | $\kappa$                |
| DESX [15,18]                     | 1            | $\kappa + n$ | $(\kappa + n)/2$               | $\max\{\kappa, n\}$     |
| double encryption [13]           | 2            | $2\kappa$    | $\kappa$                       | $\kappa$                |
| triple encryption [11,16,21]     | 3            | $3\kappa$    | $\kappa + \min\{\kappa, n\}/2$ | 90 (for 3DES)           |
| <b>double XOR-cascade [here]</b> | 2            | $\kappa + n$ | $\kappa + n/2$ (Thm. 3)        | $\kappa + n/2$ (Thm. 2) |

**Table 1.** Required number of block-cipher queries, key lengths, security lower bounds and best known attacks for various key-length extension schemes. The bounds are parameterized by the key length of the underlying block cipher (denoted by  $\kappa$ ) and its block size (denoted by  $n$ ), and are for the usual case where  $\kappa \geq n/2$ .

constructions, which are necessary to achieve higher security: Here, we prove that any construction for which distinct inputs map to distinct first queries and distinct answers from the first call imply distinct inputs to the second call admits a distinguishing attack making  $2^{\kappa+n/2}$  ideal block cipher queries and  $2^n$  construction queries. This class contains as a special case all constructions obtained by randomizing the cascade of length two using arbitrarily many key bits, including ours.

In addition, we also show (in Appendix C) that simpler randomization methods for length-two cascades admit distinguishing attacks with even lower complexity. For example, randomizing the cascade of length two as  $E_{k_2}(E_{k_1}(m \oplus z_1)) \oplus z_2$  instead of using our approach yields a simple  $2^{\max\{\kappa, n\}}$  meet-in-the-middle attack. This shows an interesting feature of our constructions, namely that while targeting CCA security (i.e., we allow for forward and backward queries to the construction), our design requires *asymmetry*, a fact which seems to contradict common wisdom.

Finally, note that all generic attacks presented in this paper (both against one-query and two-query constructions) can be mounted even if the distinguisher is only allowed to ask forward construction queries (i.e., in the CPA setting). In contrast, the construction we propose is proven to be secure even with respect to an adversary allowed to ask inverse construction queries (CCA adversary).

FINAL REMARKS. Table 1 summarizes the results of this paper in the context of previously known results. To serve as an overview, some bounds are presented in a simplified form. Note that the security of *any* key-length extension construction in our model can be upper-bounded by  $2^{\kappa+n}$  which corresponds to the trivial attack asking all possible block cipher and construction queries.

Our results and proofs are presented using Maurer’s random systems framework [22], which we review in Section 2 in a self-contained way sufficient to follow the contents of the paper.

## 2 Preliminaries

### 2.1 Basic Notation

We denote sets by calligraphic letters  $\mathcal{X}, \mathcal{Y}, \dots$ , and by  $|\mathcal{X}|, |\mathcal{Y}|, \dots$  their cardinalities. We also let  $\mathcal{X}^k$  be the set of  $k$ -tuples  $x^k = (x_1, \dots, x_k)$  of elements of  $\mathcal{X}$ . Strings are elements of  $\{0, 1\}^k$  and are usually denoted as  $s = s_1 s_2 \dots s_k$ , with  $\|$  denoting the usual string concatenation. Additionally, we let  $\text{Func}(m, \ell)$  be the set of all functions from  $\{0, 1\}^m$  to  $\{0, 1\}^\ell$  and  $\text{Perm}(n)$  be the set of all permutations of  $\{0, 1\}^n$ . In particular,  $id \in \text{Perm}(n)$  represents the identity mapping when  $n$  is understood from the context. Throughout this paper logarithms will always be to the base 2. Also,  $r_1, \dots, r_m \stackrel{\S, d}{\leftarrow} \mathcal{S}$  indicates random selection of distinct elements  $r_1, \dots, r_m$  from a set  $\mathcal{S}$ .

We denote random variables and concrete values they can take by upper-case letters  $X, Y, \dots$  and lower-case letters  $x, y, \dots$ , respectively. For events  $A$  and  $B$  and random variables  $U$  and  $V$  with ranges  $\mathcal{U}$  and  $\mathcal{V}$ , respectively, we let  $P_{U|A|V|B}$  be the corresponding conditional probability

distribution, seen as a (partial) function  $\mathcal{U} \times \mathcal{V} \rightarrow [0, 1]$ . Here the value  $P_{UA|VB}(u, v) = P[U = u \wedge A|V = v \wedge B]$  is well defined for all  $u \in \mathcal{U}$  and  $v \in \mathcal{V}$  such that  $P_{VB}(v) > 0$  and undefined otherwise. Two probability distributions  $P_U$  and  $P_{U'}$  on the same set  $\mathcal{U}$  are equal, denoted  $P_U = P_{U'}$ , if  $P_U(u) = P_{U'}(u)$  for all  $u \in \mathcal{U}$ . Conditional probability distributions are equal if the equality holds for all arguments for which both of them are defined. To emphasize the random experiment  $\mathcal{E}$  in consideration, we sometimes write it in the superscript, e.g.  $P_{U|V}^{\mathcal{E}}(u, v)$ . Finally, the complement of an event  $A$  is denoted by  $\overline{A}$ .

## 2.2 Random Systems

The presentation of this paper relies on Maurer’s random systems framework [22]. However, we stress that most of the paper remains understandable at a very high level, even without the need of a deeper understanding of the techniques behind the framework. Here we provide a self-contained introduction.

The starting point of the random-system framework is the basic observation that the input-output behavior of any kind of discrete system with inputs in  $\mathcal{X}$  and outputs in  $\mathcal{Y}$  can be described by an infinite family of functions describing, for each  $i \geq 1$ , the probability distribution of the  $i$ -th output  $Y_i \in \mathcal{Y}$  given the values of the first  $i$  inputs  $X^i \in \mathcal{X}^i$  and the previous  $i - 1$  outputs  $Y^{i-1} \in \mathcal{Y}^{i-1}$ . Formally, hence, an  $(\mathcal{X}, \mathcal{Y})$ -*(random) system*  $\mathbf{F}$  is an infinite sequence of functions  $\mathbf{p}_{Y_i|X^i Y^{i-1}}^{\mathbf{F}} : \mathcal{Y} \times \mathcal{X}^i \times \mathcal{Y}^{i-1} \rightarrow [0, 1]$  such that,  $\sum_{y_i} \mathbf{p}_{Y_i|X^i Y^{i-1}}^{\mathbf{F}}(y_i, x^i, y^{i-1}) = 1$  for all  $i \geq 1$ ,  $x^i \in \mathcal{X}^i$  and  $y^{i-1} \in \mathcal{Y}^{i-1}$ . We stress that the notation  $\mathbf{p}_{Y_i|X^i Y^{i-1}}^{\mathbf{F}}$ , by itself, involves some abuse, as we are not considering any particular random experiment with well-defined random variables  $Y_i, X^i, Y^{i-1}$  until the system will be interacting with a distinguisher (see below), in which case the random variables will exist and take the role of the transcript. In general, we shall also typically define discrete systems by a high level description, as long as the resulting conditional probability distributions can be derived uniquely from this description.

We say that a system  $\mathbf{F}$  is *deterministic* if the range of  $\mathbf{p}_{Y_i|X^i Y^{i-1}}^{\mathbf{F}}$  is  $\{0, 1\}$  for all  $i \geq 1$ . Moreover, it is *stateless* if the probability distribution of each output depends only on the current input, i.e., if there exists a distribution  $\mathbf{p}_{Y|X} : \mathcal{Y} \times \mathcal{X} \rightarrow [0, 1]$  such that  $\mathbf{p}_{Y_i|X^i Y^{i-1}}^{\mathbf{F}}(y_i, x^i, y^{i-1}) = \mathbf{p}_{Y|X}(y_i, x_i)$  for all  $y_i, x^i$  and  $y^{i-1}$ .

We also consider systems  $\mathbf{C}^{\mathbf{F}}$  that arise from *constructions*  $\mathbf{C}^{(\cdot)}$  accessing a sub-system  $\mathbf{F}$ . Note that while a construction  $\mathbf{C}^{(\cdot)}$  does not define a random system by itself,  $\mathbf{C}^{\mathbf{F}}$  does define a random system. The notions of being deterministic and of being stateless naturally extend to constructions.<sup>9</sup> We also consider the *parallel composition* of two (possibly dependent) discrete systems  $\mathbf{F}$  and  $\mathbf{G}$ , denoted  $(\mathbf{F}, \mathbf{G})$ , which is the system that allows queries to both systems  $\mathbf{F}$  and  $\mathbf{G}$ .

EXAMPLES. A *random function*  $\mathbf{F} : \{0, 1\}^m \rightarrow \{0, 1\}^n$  is a system which implements a function  $f$  initially chosen according to some distribution on the set  $\text{Func}(m, n)$ .<sup>10</sup> In particular, the *uniform random function (URF)*  $\mathbf{R} : \{0, 1\}^m \rightarrow \{0, 1\}^n$  realizes a uniformly chosen function  $f \in \text{Func}(m, n)$ , whereas the *uniform random permutation (URP)* on  $\{0, 1\}^n$ , denoted  $\mathbf{P} : \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$ , realizes a uniformly chosen permutation  $P \in \text{Perm}(n)$  allowing both forward queries of the form  $(x, +)$  returning  $P(x)$  as well as backward queries  $(y, -)$  returning  $P^{-1}(y)$ . More generally, we meet the convention (for the purpose of this paper) that any system realizing a random function (possibly

<sup>9</sup> We dispense with a formal definition. However, we point out that we allow a stateless construction to keep a state during invocations of its subsystem.

<sup>10</sup> As is the case with the notion of a random variable, the word “random” does not imply uniformity of the distribution.

by means of a construction) which is a permutation will *always* allow both forward and backward queries.

Another important example of a random function is the *ideal block cipher*  $\mathbf{E} : \{0, 1\}^\kappa \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  which realizes an independent uniform random permutation  $\mathbf{E}_k \in \text{Perm}(n)$  for each key  $k \in \{0, 1\}^\kappa$ ; in particular, the system allows both forward and backward queries to each  $\mathbf{E}_k$ .

Finally, note that with some abuse of notation, we often write  $\mathbf{E}_k$  or  $\mathbf{P}$  to refer to the randomly chosen permutation  $P$  implemented by the system  $\mathbf{E}_k$  or  $\mathbf{P}$ , respectively.

**DISTINGUISHERS AND INDISTINGUISHABILITY.** A *distinguisher*  $\mathbf{D}$  for an  $(\mathcal{X}, \mathcal{Y})$ -random system asking  $q$  queries is a  $(\mathcal{Y}, \mathcal{X})$ -random system which is “one query ahead:” its input-output behavior is defined by the conditional probability distributions of its queries  $\mathbf{p}_{X_i|X^{i-1}Y^{i-1}}^{\mathbf{D}}$  for all  $1 \leq i \leq q$ . (The first query of  $\mathbf{D}$  is determined by  $\mathbf{p}_{X_1}^{\mathbf{D}}$ .) After the distinguisher asks all  $q$  queries, it outputs a bit  $W_q$  depending on the transcript  $(X^q, Y^q)$ . For a random system  $\mathbf{F}$  and a distinguisher  $\mathbf{D}$ , let  $\mathbf{DF}$  be the random experiment where  $\mathbf{D}$  interacts with  $\mathbf{F}$ , with the distributions of the transcript  $(X^q, Y^q)$  and of the bit  $W_q$  being uniquely defined by their conditional probability distributions. Then, for two  $(\mathcal{X}, \mathcal{Y})$ -random systems  $\mathbf{F}$  and  $\mathbf{G}$ , the *distinguishing advantage* of  $\mathbf{D}$  in distinguishing systems  $\mathbf{F}$  and  $\mathbf{G}$  by  $q$  queries is the quantity  $\Delta^{\mathbf{D}}(\mathbf{F}, \mathbf{G}) = |\mathbf{P}^{\mathbf{DF}}(W_q = 1) - \mathbf{P}^{\mathbf{DG}}(W_q = 1)|$ . We are usually interested in the maximal distinguishing advantage over all distinguishers asking  $q$  queries, which we denote by  $\Delta_q(\mathbf{F}, \mathbf{G}) = \max_{\mathbf{D}} \Delta^{\mathbf{D}}(\mathbf{F}, \mathbf{G})$  (with  $\mathbf{D}$  ranging over all such distinguishers).

For a random system  $\mathbf{F}$ , we often consider an internal *monotone condition* defined on it. Such a condition is initially satisfied (true), but once it gets violated, it cannot become true again (hence the name monotone). Typically, the condition captures whether the behavior of the system meets some additional requirement (e.g. distinct outputs, consistent outputs) or this was already violated during the interaction. We formalize such a condition by a sequence of events  $\mathcal{A} = A_0, A_1, \dots$  such that  $A_0$  always holds, and  $A_i$  holds if the condition holds after query  $i$ . The probability that a distinguisher  $\mathbf{D}$  issuing  $q$  queries to  $\mathbf{F}$  makes a monotone condition  $\mathcal{A}$  fail in the random experiment  $\mathbf{DF}$  is denoted by  $\nu^{\mathbf{D}}(\mathbf{F}, \overline{A}_q) = \mathbf{P}^{\mathbf{DF}}(\overline{A}_q)$  and we are again interested in the maximum over all such distinguishers, denoted by  $\nu(\mathbf{F}, \overline{A}_q) = \max_{\mathbf{D}} \nu^{\mathbf{D}}(\mathbf{F}, \overline{A}_q)$ . For a random system  $\mathbf{F}$  with a monotone condition  $\mathcal{A} = A_0, A_1, \dots$  and a random system  $\mathbf{G}$ , we say that  $\mathbf{F}$  *conditioned on*  $\mathcal{A}$  *is equivalent to*  $\mathbf{G}$ , denoted  $\mathbf{F}|\mathcal{A} \equiv \mathbf{G}$ , if  $\mathbf{p}_{Y_i|X^iY^{i-1}A_i}^{\mathbf{F}} = \mathbf{p}_{Y_i|X^iY^{i-1}}^{\mathbf{G}}$  for  $i \geq 1$ , for all arguments for which  $\mathbf{p}_{Y_i|X^iY^{i-1}A_i}^{\mathbf{F}}$  is defined. Intuitively, this captures the fact that as long as the condition  $\mathcal{A}$  holds in  $\mathbf{F}$ , it behaves the same as  $\mathbf{G}$ .

Let  $\mathbf{F}$  be a random system with a monotone condition  $\mathcal{A}$ . Following [25], we define  $\mathbf{F}$  *blocked by*  $\mathcal{A}$  to be a new random system that behaves exactly like  $\mathbf{F}$  while the condition  $\mathcal{A}$  is satisfied. Once  $\mathcal{A}$  is violated, it only outputs a special blocking symbol  $\perp$  not contained in the output alphabet of  $\mathbf{F}$ .

We make use of the following helpful claims proven in previous papers. Below, we also present an informal explanation of their merits.

**Lemma 1.** *Let  $\mathbf{C}^{(\cdot)}$  and  $\mathbf{C}'^{(\cdot)}$  be two constructions invoking a subsystem, and let  $\mathbf{F}$  and  $\mathbf{G}$  be random systems. Let  $\mathcal{A}$  and  $\mathcal{B}$  be two monotone conditions defined on  $\mathbf{F}$  and  $\mathbf{G}$ , respectively.*

- (i) [22, Theorem 1] *If  $\mathbf{F}|\mathcal{A} \equiv \mathbf{G}$  then  $\Delta_q(\mathbf{F}, \mathbf{G}) \leq \nu(\mathbf{F}, \overline{A}_q)$ .*
- (ii) [16, Lemma 2] *Let  $\mathbf{F}^\perp$  denote the random system  $\mathbf{F}$  blocked by  $\mathcal{A}$  and let  $\mathbf{G}^\perp$  denote  $\mathbf{G}$  blocked by  $\mathcal{B}$ . Then for every distinguisher  $\mathbf{D}$  asking  $q$  queries we have  $\Delta^{\mathbf{D}}(\mathbf{F}, \mathbf{G}) \leq \Delta_q(\mathbf{F}^\perp, \mathbf{G}^\perp) + \nu^{\mathbf{D}}(\mathbf{F}, \overline{A}_q)$ .*
- (iii) [22, Lemma 5]  *$\Delta_q(\mathbf{C}^{\mathbf{F}}, \mathbf{C}^{\mathbf{G}}) \leq \Delta_{q'}(\mathbf{F}, \mathbf{G})$ , where  $q'$  is the maximum number of invocations of any internal system  $\mathbf{H}$  for any sequence of  $q$  queries to  $\mathbf{C}^{\mathbf{H}}$ , if such a value is defined.*

(iv) [16, Lemma 3] There exists a fixed permutation  $S \in \text{Perm}(n)$  (represented by a deterministic stateless system) such that  $\Delta_q(\mathbf{C}^{\mathbf{P}}, \mathbf{C}'^{\mathbf{P}}) \leq \Delta_q(\mathbf{C}^{\mathbf{S}}, \mathbf{C}'^{\mathbf{S}})$ .

The first claim can be seen as a generalized version of the Fundamental Lemma of Game-Playing for the context of random systems, stating that if two systems are equivalent as long as some condition is satisfied, then the advantage in distinguishing these systems can be upper-bounded by the probability of violating this condition. The second claim is even more general, analyzing the situation where the systems are not equivalent even if the conditions defined on them are satisfied, but their behavior is similar (which is captured by the term  $\Delta_q(\mathbf{F}^\perp, \mathbf{G}^\perp)$ ). The third claim states the intuitive fact that interacting with the distinguished systems through an additional enveloping construction  $\mathbf{C}$  cannot improve the distinguishing advantage and the last claim is just an averaging argument over all the possible values taken by  $\mathbf{P}$ .

### 3 Generic Attacks against Efficient Key-Length Extension Schemes

We start by addressing the following question: *What is the maximum achievable security level for very efficient key-length extension schemes?* To this end, this section presents generic distinguishing attacks against one- and two-call block-cipher constructions in Sections 3.1 and 3.2, respectively. These attacks are in the same spirit as the recent line of work on generic attacks on hash functions (cf. e.g. [27,29,30]). Along the same lines, here attack complexity will be measured in terms of query- rather than time-complexity. This allows us to consider arbitrary constructions, while being fully sufficient to assess security in the ideal cipher model, where distinguishers are computationally unrestricted.

More formally, we consider stateless and deterministic (keyed) constructions  $\mathbf{C}^{(\cdot)}$  invoking an ideal cipher  $\mathbf{E} : \{0, 1\}^\kappa \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  to implement a function  $\mathbf{C}^{\mathbf{E}} : \{0, 1\}^{\kappa'} \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  to serve as a block cipher with key length  $\kappa'$ . We assume that the construction  $\mathbf{C}^{\mathbf{E}}$  realizes a permutation for each  $k' \in \{0, 1\}^{\kappa'}$  and hence it also provides the interface for inverse queries as indicated. Consequently, for a random (secret)  $\kappa'$ -bit string  $K'$ , we let  $\mathbf{C}_{K'}^{\mathbf{E}}$  denote the system which only gives access to the permutation  $\mathbf{C}^{\mathbf{E}}(K', \cdot)$  and its inverse (i.e., takes inputs from  $\{0, 1\}^n \times \{+, -\}$ ). (In fact, none of the attacks in this section will require backward queries.)

#### 3.1 One-Query Constructions

Throughout this section, we assume that  $\mathbf{C}^{(\cdot)}$ , to evaluate input  $(x, +)$  for  $x \in \{0, 1\}^n$  under a key  $k' \in \{0, 1\}^{\kappa'}$ , makes exactly one query to the underlying subsystem, and we denote this query as  $q(k', x)$ . We consider two different cases, depending on the structure of  $q(\cdot, \cdot)$ , before deriving the final attack. Note that we omit the proofs of our claims in this section and only provide some intuition. All statements are proved (partially in a more general setting) in Appendix A.

THE INJECTIVE CASE. We first consider the case where the mapping  $x \mapsto q(k', x)$  is injective for each  $k'$ . We shall denote this as a *one-injective-query construction*. In this case, distinct queries to  $\mathbf{C}_{k'}^{\mathbf{E}}$  lead to distinct internal queries to  $\mathbf{E}$  and hence if the distinguisher queries both  $\mathbf{C}_{K'}^{\mathbf{E}}$  and  $\mathbf{E}$  at sufficiently many random positions, one can expect that during the evaluation of the outer queries,  $\mathbf{C}_{K'}^{(\cdot)}$  asks  $\mathbf{E}$  for a value that was also asked by the distinguisher. If this occurs, the distinguisher can, while trying all possible keys  $k'$ , evaluate  $\mathbf{C}_{k'}^{(\cdot)}$  on its own by simulating  $\mathbf{C}^{(\cdot)}$  and using the response from  $\mathbf{E}$ ; and by comparing the outcomes it can distinguish the construction from a truly random permutation. This is the main idea behind the following lemma.

**Lemma 2.** Let  $\mathbf{E}: \{0, 1\}^\kappa \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  be an ideal block cipher, let  $\mathbf{C}^{(\cdot)}: \{0, 1\}^{\kappa'} \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  be a one-injective-query construction and let  $\mathbf{P}$  be a URP on  $\{0, 1\}^n$ . Then, for a random key  $K' \in \{0, 1\}^{\kappa'}$  and every parameter  $0 < t < 2^{\min\{n, \kappa'\}-1}$ ,<sup>11</sup> there exists a distinguisher  $\mathbf{D}$  such that

$$\Delta^{\mathbf{D}}((\mathbf{E}, \mathbf{C}_{K'}^{\mathbf{E}}), (\mathbf{E}, \mathbf{P})) \geq 1 - 2/t - 2^{\kappa' - t \cdot (n-1)},$$

and which makes at most  $4t \cdot 2^{\max\{(\kappa+n)/2, \kappa\}}$  queries to the block cipher  $\mathbf{E}$ , as well as at most  $2 \cdot 2^{\min\{(\kappa+n)/2, n\}}$  forward queries to either of  $\mathbf{C}_{K'}^{\mathbf{E}}$  and  $\mathbf{P}$ .

The above lemma covers most of the natural one-query constructions, since these typically satisfy the injectivity requirement (e.g. the DESX construction). In the following we see that constructions asking non-injective queries do not achieve any improvement in security.

NON-INJECTIVE QUERIES. We now permit that the construction  $\mathbf{C}^{(\cdot)}$  might, for some key  $k'$ , invoke the underlying ideal cipher in a *non-injective* way, i.e.,  $q(k', \cdot)$  is not an injective map. We prove that, roughly speaking, such a construction  $\mathbf{C}_{K'}^{\mathbf{E}}$  might be distinguishable from a URP  $\mathbf{P}$  based solely on an entropy argument. The intuitive reasoning is that if  $\mathbf{C}^{(\cdot)}$  allows on average (over the choice of the key  $k'$ ) that too many queries  $x$  map to the same  $q(k', x)$ , then it also does not manage to obtain sufficient amount of randomness from the underlying random function to simulate  $\mathbf{P}$  convincingly, opening the door to a distinguishing attack. In the following, let  $q(k') = |\{q(k', x) : x \in \{0, 1\}^n\}|$  for all  $k' \in \{0, 1\}^{\kappa'}$ .

**Lemma 3.** Let  $\mathbf{C}^{(\cdot)}: \{0, 1\}^{\kappa'} \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  be a one-query construction, let  $\mathbf{P}$  be a URP on  $\{0, 1\}^n$  and let  $\mathbf{E}: \{0, 1\}^\kappa \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  be an ideal block cipher. Also, let  $K' \in \{0, 1\}^{\kappa'}$  be a random key, and assume that there exists  $q^*$  such that  $q(K') \leq q^*$  with probability at least  $\frac{1}{2}$ . Then, there exists a distinguisher  $\mathbf{D}$  asking  $2^n$  forward queries such that

$$\Delta^{\mathbf{D}}(\mathbf{C}_{K'}^{\mathbf{E}}, \mathbf{P}) \geq \frac{1}{2} - 2^{\kappa' + n \cdot q^* - \log(2^n!)}.$$

PUTTING THE PIECES TOGETHER. We can combine the techniques used to prove Lemma 2 (somewhat relaxing the injectivity requirement) and Lemma 3, to obtain the following final theorem yielding an attack for arbitrary one-query block-cipher constructions.

**Theorem 1.** Let  $n \geq 6$  and  $\kappa' \leq 2^n - 1$ , let  $\mathbf{E}: \{0, 1\}^\kappa \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  be an ideal block cipher, let  $\mathbf{C}^{(\cdot)}: \{0, 1\}^{\kappa'} \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  be a one-query construction, and let  $\mathbf{P}$  be a URP on  $\{0, 1\}^n$ . Then, for a random key  $K' \in \{0, 1\}^{\kappa'}$  and for all parameters  $0 < t < 2^{n-2}$ , there exists a distinguisher  $\mathbf{D}$  such that

$$\Delta^{\mathbf{D}}((\mathbf{E}, \mathbf{C}_{K'}^{\mathbf{E}}), (\mathbf{E}, \mathbf{P})) \geq \min\left\{\frac{1}{4}, \frac{1}{2} - \frac{2}{t} - 2^{\kappa' - t \cdot (n-1)}\right\},$$

and which asks at most  $8t \cdot 2^\kappa$  queries to  $\mathbf{E}$  and  $2^n$  forward queries to either of  $\mathbf{C}_{K'}^{\mathbf{E}}$  and  $\mathbf{P}$ .

Theorem 1 shows that no one-query construction can achieve security beyond  $2^{\max\{\kappa, n\}}$  queries, hence in our search for efficient key-length extension schemes we have to turn our attention towards constructions issuing at least two queries.

<sup>11</sup> Roughly speaking, higher  $t$  increases the advantage but also the required number of queries; we obtain the desired bound using a constant  $t$ . For a first impression, consider e.g.  $t = 4$  and  $\kappa' \approx 2n$ .

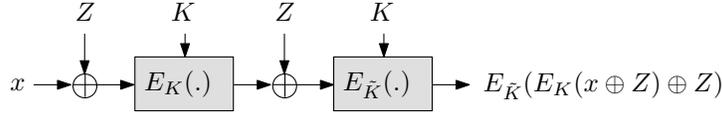


Fig. 1. The double XOR-cascade construction analyzed in Theorem 3.

### 3.2 Two-Query Constructions

We now consider an arbitrary deterministic stateless construction  $\mathbf{C}^{(\cdot)} : \{0, 1\}^{\kappa'} \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  that makes exactly two queries to an ideal block cipher  $\mathbf{E} : \{0, 1\}^{\kappa} \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  to evaluate each query. In the following, these constructions shall be referred to as *two-query constructions*. We denote by  $q_1(k', x) \in \{0, 1\}^{\kappa} \times \{0, 1\}^n \times \{+, -\}$  the first query  $\mathbf{C}^{(\cdot)}$  asks its subsystem when it is itself being asked a forward query  $(k', x, +)$ . Moreover, we denote by  $q_2(k', x, s) \in \{0, 1\}^{\kappa} \times \{0, 1\}^n \times \{+, -\}$  the second query it asks when it is itself being asked a forward query  $(k', x, +)$  and the answer to the first query  $q_1(k', x)$  was  $s \in \{0, 1\}^n$ . Since  $\mathbf{C}^{(\cdot)}$  is deterministic and stateless, both  $q_1$  and  $q_2$  are well-defined mappings.

**Theorem 2.** *Let  $\mathbf{C}^{(\cdot)} : \{0, 1\}^{\kappa'} \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  be a two-query construction satisfying the following two conditions:*

1. *for every  $k' \in \{0, 1\}^{\kappa'}$  the mapping  $q_1(k', \cdot)$  is injective,*
2. *distinct answers to the first query imply distinct second queries, i.e., for every  $k' \in \{0, 1\}^{\kappa'}$  and every  $x, x' \in \{0, 1\}^n$  if  $s \neq s'$  then  $q_2(k', x, s) \neq q_2(k', x', s')$ .*

*Then for a random key  $K' \in \{0, 1\}^{\kappa'}$ , for a URP  $\mathbf{P}$  on  $\{0, 1\}^n$  and for every parameter  $0 < t < 2^{n/2-1}$ , there exists a distinguisher  $\mathbf{D}$  such that*

$$\Delta^{\mathbf{D}}((\mathbf{E}, \mathbf{C}_{K'}^{\mathbf{E}}), (\mathbf{E}, \mathbf{P})) \geq 1 - 2/t - 13 \cdot 2^{-\frac{n}{2}} - 2^{\kappa'-t \cdot (n-1)},$$

*where  $\mathbf{D}$  makes at most  $2(t+4) \cdot 2^{\kappa+n/2}$  queries to  $\mathbf{E}$  as well as  $2^n$  forward queries to either of  $\mathbf{C}_{K'}^{\mathbf{E}}$  and  $\mathbf{P}$ .*

The proof of Theorem 2 is given in Appendix B.

Hence, no two-query construction from the large class described in the above theorem can achieve security beyond  $2^{\kappa+n/2}$  queries. In the following section we present a simple and efficient construction from this class that achieves the above limit.

## 4 The Double XOR-Cascade Construction

We present a two-query construction matching the upper bound  $2^{\kappa+n/2}$  on security proved in the previous section. The construction, which we call the *double XOR-cascade construction* (2XOR), consists of two applications of the block-cipher interleaved with two XOR operations: Given a  $(\kappa, n)$ -block cipher  $E$ , we define the  $(\kappa + n, n)$ -block cipher  $2\text{XOR}^E$  such that

$$2\text{XOR}_{k,z}^E(m) = E_{\tilde{k}}(E_k(m \oplus z) \oplus z)$$

for all  $k \in \{0, 1\}^{\kappa}$ ,  $z, m \in \{0, 1\}^n$ , and where  $\tilde{k} = \pi(k)$  for some understood fixpoint-free permutation  $\pi \in \text{Perm}(\kappa)$  (e.g.,  $\pi(k) = k \oplus 0^{\kappa-1}1$ , i.e.,  $\pi$  flips the last bit). The construction is depicted in Figure 1. Note that both XOR transformations use the same value  $z$  and the two block-cipher calls use two distinct keys such that one can be deterministically derived from the other one. We also consider a construction  $2\text{XOR}'$  of a  $(2\kappa + n)$ -block cipher where  $\tilde{k}$  is replaced by an (independent)  $\kappa$ -bit key.

SECURITY OF 2XOR. We now discuss the security of the double XOR-cascade construction in the ideal cipher model. To this end, let  $\mathbf{X}^{(\cdot)}: \{0, 1\}^\kappa \times \{0, 1\}^n \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  denote a (deterministic stateless) construction which expects a subsystem  $\mathbf{E}: \{0, 1\}^\kappa \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  realizing a block cipher.  $\mathbf{X}^{\mathbf{E}}$  then answers each query  $(k, z, x, +)$  by  $\mathbf{E}_k(\mathbf{E}_k(x \oplus z) \oplus z)$  and each query  $(k, z, y, -)$  by  $\mathbf{E}_k^{-1}(\mathbf{E}_k^{-1}(y) \oplus z) \oplus z$ . As before, for randomly chosen (secret) keys  $(K, Z) \in \{0, 1\}^\kappa \times \{0, 1\}^n$ , we let  $\mathbf{X}_{K,Z}^{\mathbf{E}}$  be the system which gives access to the permutation  $\mathbf{X}^{\mathbf{E}}(K, Z, \cdot)$  in both directions (i.e., takes inputs from  $\{0, 1\}^n \times \{+, -\}$ ).

**Theorem 3.** *Let  $\mathbf{P}$  and  $\mathbf{E}$  denote a URP on  $\{0, 1\}^n$  and an ideal  $(\kappa, n)$ -block cipher, respectively; let  $(K, Z) \in \{0, 1\}^\kappa \times \{0, 1\}^n$  be uniformly chosen keys. For the construction  $\mathbf{X}_{K,Z}^{(\cdot)}$  defined as above, and for every distinguisher  $\mathbf{D}$  making  $q$  queries to  $\mathbf{E}$ ,*

$$\Delta^{\mathbf{D}}((\mathbf{E}, \mathbf{X}_{K,Z}^{\mathbf{E}}), (\mathbf{E}, \mathbf{P})) \leq 4 \cdot \left( \frac{q}{2^{\kappa+n/2}} \right)^{2/3}.$$

*In particular,  $\mathbf{D}$  can make arbitrarily many queries to either of  $\mathbf{X}_{K,Z}^{\mathbf{E}}$  and  $\mathbf{P}$ .*

We also note that an analogous statement for the construction 2XOR' can be easily derived from the presented claim.

PROOF INTUITION. The proof, given below, follows a two-step approach. In the first part, we prove that for any parameter  $h \leq 2^{n/2}$ , the above advantage is upper bounded by  $\varepsilon(h) + \frac{q}{h2^{\kappa-1}}$ , where  $\varepsilon(h)$  is an upper bound on the advantage of a  $h$ -query distinguisher in telling apart the following two settings, in both of which it issues both forward and backward queries to two permutations  $\pi_1, \pi_2 \in \text{Perm}(n)$ :

1. In the first case,  $\pi_1, \pi_2$  are chosen uniformly and independently.
2. In the second setting, a uniform  $n$ -bit string  $Z$  is chosen, and  $\pi_1$  and  $\pi_2$  are chosen uniformly at random such that  $\pi_2(\pi_1(\cdot \oplus Z) \oplus Z) = \text{id}$ .

This step follows a pattern similar to the one used in [11,16] to analyze the security of plain cascades, but with obvious modifications and extra care to take into account randomization as well as key dependency.

Then, the main technical part of the proof consists of proving a bound  $3h^2/2^{n+1}$  on  $\varepsilon(h)$ , which is a new result of independent interest. The intuition here is that without knowing  $Z$ , it is hard to come up with two queries, one to  $\pi_1$  and one to  $\pi_2$ , which result in input-output pairs  $\pi_1(x) = y$  and  $\pi_2(x') = y'$  satisfying  $x = y' \oplus Z$  and  $x' = y \oplus Z$  simultaneously. However, as long as this does not happen, both permutations appear independent and random.

We stress that our double-randomization is crucial here: omitting one of the randomization steps, as well as adding a third randomization step for the same  $Z$ , would all result in invalidating the argument. In particular, Appendix C provides some useful extra intuition as for why other simpler randomization methods for the cascade fail to provide the required security level.

*Proof.* We start by noting that the system  $(\mathbf{E}, \mathbf{X}_{K,Z}^{\mathbf{E}})$  simply can be seen as providing an interface to query  $2^\kappa + 1$  (dependent) permutations

$$\mathbf{E}_{k_1}, \mathbf{E}_{k_2}, \dots, \mathbf{E}_{k_{2^\kappa}}, \mathbf{E}_{\tilde{K}}(\mathbf{E}_K(\cdot \oplus Z) \oplus Z),$$

each both in forward and backward direction, where  $k_1, k_2, \dots, k_{2^\kappa}$  is an enumeration of the  $\kappa$ -bit strings. By the group structure of  $\text{Perm}(n)$  under composition, the joint distribution of these permutations does not change if we start by choosing the last permutation uniformly at random,

i.e., we replace it by  $\mathbf{P}$ , then choose  $K$  and  $Z$  and finally choose the permutations of the block cipher independently and randomly except for the one corresponding to the key  $\tilde{K}$ , which we set to  $x \mapsto \mathbf{P}(\mathbf{E}_K^{-1}(x \oplus Z) \oplus Z)$ . Hence, let  $\mathbf{G}^{(\cdot)}$  be a system that expects a single permutation as its subsystem (let us denote it  $P$ ) and itself provides an interface to a block cipher (let us denote it  $G$ ). It answers queries to  $G$  in the following way: in advance, it chooses random keys  $(K, Z)$  and then generates random independent permutations for  $G$  used with any key except  $\tilde{K}$ . For  $\tilde{K}$ ,  $\mathbf{G}$  realizes the permutation  $x \mapsto P(G_{\tilde{K}}^{-1}(x \oplus Z) \oplus Z)$ , querying  $P$  for any necessary values. Then the above argument shows that  $(\mathbf{E}, \mathbf{X}_{K,Z}^{\mathbf{E}}) = (\mathbf{G}^{\mathbf{P}}, \mathbf{P})$  and hence we obtain

$$\Delta_q((\mathbf{E}, \mathbf{X}_{K,Z}^{\mathbf{E}}), (\mathbf{E}, \mathbf{P})) = \Delta_q((\mathbf{G}^{\mathbf{P}}, \mathbf{P}), (\mathbf{E}, \mathbf{P})) \leq \Delta_q((\mathbf{G}^S, S), (\mathbf{E}, S)),$$

where the last inequality follows from claim (iv) in Lemma 1 and  $S$  denotes the fixed permutation whose existence is guaranteed by this claim. Since  $S$  is fixed and hence known to the distinguisher, it makes no sense to query it and thus it remains to bound  $\Delta_q(\mathbf{G}^S, \mathbf{E})$  for any permutation  $S$ . From now on, we denote the system  $\mathbf{G}^S$  by  $\mathbf{G}$  to simplify the notation.

We shall refer to all forward or backwards queries to  $G$  involving the permutations indexed by  $K$  or  $\tilde{K}$  as *relevant*. Similarly, the system  $\mathbf{E}$  can be seen as also choosing some random key  $K$  (and hence  $\tilde{K}$ ), this just does not affect its behavior, and we can hence define relevant queries for  $\mathbf{E}$  in an analogous way. Let  $\mathcal{A}^h$  and  $\mathcal{B}^h$  denote monotone conditions defined on systems  $\mathbf{E}$  and  $\mathbf{G}$  respectively, such that each of these conditions remains satisfied as long as at most  $h$  of the queries asked so far were relevant. The parameter  $h$  will be chosen optimally at the end of the proof. We require  $h < 2^{n/2}$ .

It is easy to upper-bound the probability of asking more than  $h$  relevant queries in  $\mathbf{E}$ : since the key  $K$  does not affect the responses of the system (and therefore the behavior is also independent of the associated monotone condition), we only have to consider non-adaptive strategies. Hence, for any distinguisher  $\mathbf{D}$  asking  $q$  queries, the expected number of relevant queries among them is  $q \cdot 2^{1-\kappa}$  and using Markov inequality, we obtain  $\nu(\mathbf{E}, \overline{\mathcal{A}}_q^h) \leq q/h2^{\kappa-1}$ . Let  $\mathbf{E}^\perp$  and  $\mathbf{G}^\perp$  denote the systems  $\mathbf{E}$  and  $\mathbf{G}$  blocked by  $\mathcal{A}^h$  and  $\mathcal{B}^h$ , respectively. Then we can apply claim (ii) of Lemma 1 to obtain

$$\Delta_q(\mathbf{G}, \mathbf{E}) \leq \Delta_q(\mathbf{G}^\perp, \mathbf{E}^\perp) + \nu(\mathbf{E}, \overline{\mathcal{A}}_q^h) \leq \Delta_q(\mathbf{G}^\perp, \mathbf{E}^\perp) + q/h2^{\kappa-1}.$$

Now, one can observe that the systems  $\mathbf{G}^\perp$  and  $\mathbf{E}^\perp$  only differ in a small part. More specifically, we have  $\mathbf{G}^\perp = \mathbf{C}^{\mathbf{S}}$  and  $\mathbf{E}^\perp = \mathbf{C}^{\mathbf{T}}$ , where:

- $\mathbf{S}$  is a system that chooses  $Z \in \{0, 1\}^n$  at random and provides access (by means of both forward and backward queries) to two randomly chosen permutations  $\pi_1, \pi_2$  on  $\{0, 1\}^n$  such that they satisfy the equation  $\pi_2(\pi_1(\cdot \oplus Z) \oplus Z) = id$ ;
- $\mathbf{T}$  is a system providing access (by means of both forward and backward queries) to two uniformly random permutations  $\pi_1, \pi_2 \in \text{Perm}(n)$  that are independent;
- $\mathbf{C}^{(\cdot)}$  is a (randomized) construction that expects a subsystem which provides two permutations  $\pi_1$  and  $\pi_2$ .  $\mathbf{C}^{(\cdot)}$  itself provides access to a block cipher  $C$  as follows: it chooses a uniformly random key  $K$  and sets  $C_K := \pi_1$  and  $C_{\tilde{K}} := \pi_2 \circ S$ . ( $\mathbf{C}$  only queries its subsystem once it is asked a relevant query). The permutations for all other keys are chosen independently at random. Moreover,  $\mathbf{C}$  only allows  $h$  relevant queries, after that it returns  $\perp$ .

By Lemma 1(iii), the above observation gives us  $\Delta_q(\mathbf{G}^\perp, \mathbf{E}^\perp) \leq \Delta_h(\mathbf{S}, \mathbf{T})$ .

Hence it remains to bound  $\Delta_h(\mathbf{S}, \mathbf{T})$ . We start by taking a different view of the internal workings of the system  $\mathbf{S}$ . Once the values  $Z, \pi_1, \pi_2$  are chosen, the internal state of  $\mathbf{S}$  can be represented by

a set  $\mathcal{T}$  of  $2^n$  4-tuples  $(x_1, y_1, x_2, y_2)$  such that  $\pi_1(x_1) = y_1$  and  $\pi_2(x_2) = y_2$ , and  $x_2 = y_1 \oplus Z$  and  $x_1 = y_2 \oplus Z$ . For any  $\mathcal{I} \subseteq \{1, \dots, 4\}$ , let  $\mathcal{T}_{\mathcal{I}}$  be the projection of  $\mathcal{T}$  on the components in  $\mathcal{I}$ . Then note that for any two distinct tuples  $(x_1, y_1, x_2, y_2), (x'_1, y'_1, x'_2, y'_2) \in \mathcal{T}$  we have  $x_1 \neq x'_1, y_1 \neq y'_1, x_2 \neq x'_2$ , and  $y_2 \neq y'_2$ , in other words  $\mathcal{T}_{\{i\}} = \{0, 1\}^n$  for every  $i \in \{1, \dots, 4\}$ .

Equivalently, it is not hard to verify that  $\mathbf{S}$  can be implemented using lazy-sampling to set up  $\mathcal{T}$ : Initially,  $\mathcal{T} = \emptyset$  and  $Z$  is a uniform  $n$ -bit string. Then,  $\mathbf{S}$  answers queries as follows:

- Upon a query  $\pi_1(x)$ , it returns  $y$  if  $(x, y) \in \mathcal{T}_{\{1,2\}}$  for some  $y$ . Otherwise, it returns a random  $y \in \{0, 1\}^n \setminus \mathcal{T}_{\{2\}}$  and adds  $(x, y, y \oplus Z, x \oplus Z)$  to  $\mathcal{T}$ .
- Upon a query  $\pi_1^{-1}(y)$ , it returns  $x$  if  $(x, y) \in \mathcal{T}_{\{1,2\}}$  for some  $x$ . Otherwise, it returns a random  $x \in \{0, 1\}^n \setminus \mathcal{T}_{\{1\}}$  and adds  $(x, y, y \oplus Z, x \oplus Z)$  to  $\mathcal{T}$ .
- Upon a query  $\pi_2(x)$ , it returns  $y$  if  $(x, y) \in \mathcal{T}_{\{3,4\}}$  for some  $y$ . Otherwise, it returns a random  $y \in \{0, 1\}^n \setminus \mathcal{T}_{\{4\}}$  and adds  $(y \oplus Z, x \oplus Z, x, y)$  to  $\mathcal{T}$ .
- Upon a query  $\pi_2^{-1}(y)$ , it returns  $x$  if  $(x, y) \in \mathcal{T}_{\{3,4\}}$  for some  $x$ . Otherwise, it returns a random  $x \in \{0, 1\}^n \setminus \mathcal{T}_{\{3\}}$  and adds  $(y \oplus Z, x \oplus Z, x, y)$  to  $\mathcal{T}$ .

We consider an intermediate system  $\mathbf{S}'$  obtained from  $\mathbf{S}$ : In addition to  $\mathcal{T}$ , it also keeps track of sets  $\mathcal{P}_1$  and  $\mathcal{P}_2$ , both consisting of ordered pairs of  $n$ -bit strings. (Again  $\mathcal{P}_{i,1}$  and  $\mathcal{P}_{i,2}$  denote the strings appearing as first and second component in  $\mathcal{P}_i$ , respectively.) Initially each  $\mathcal{P}_i$  is empty and during the experiment,  $\mathcal{P}_i$  keeps track of input-output pairs for  $\pi_i$  which were already defined by directly answering a  $\pi_i$  query in either direction (as opposed to those that were defined internally by  $\mathbf{S}'$  when answering a  $\pi_{3-i}$  query). Concretely,  $\mathbf{S}'$  answers a query  $\pi_1(x)$  by  $y$  if  $(x, y) \in \mathcal{T}_{\{1,2\}} \cup \mathcal{P}_1$  for some  $y$ . Otherwise, it returns a uniformly chosen  $y \in \{0, 1\}^n \setminus \mathcal{P}_{1,2}$  and adds  $(x, y)$  to  $\mathcal{P}_1$ . Moreover, if  $y \notin \mathcal{T}_{\{2\}}$ , it also adds the tuple  $(x, y, y \oplus Z, x \oplus Z)$  to  $\mathcal{T}$ . Queries  $\pi_1^{-1}(y)$ ,  $\pi_2(x)$ , and  $\pi_2^{-1}(y)$  are answered in a symmetric fashion. Having this description of  $\mathbf{S}'$ , note that we obtain the system  $\mathbf{T}$  if a query  $\pi_1(x)$  is answered by some given  $y$  only if  $(x, y) \in \mathcal{P}_1$ , and otherwise a fresh random output is generated (but the 4-tuples are still added to  $\mathcal{T}$  as above).

We now define two monotone conditions  $\mathcal{A}$  and  $\mathcal{B}$  on  $\mathbf{S}'$ :

- The condition  $\mathcal{A} = A_0, A_1, \dots$  fails at the first query  $\pi_i(x)$  answered by a random  $y$  which satisfies  $y \in \mathcal{T}_{\{2(i-1)+2\}}$ , or  $\pi_i^{-1}(y)$  answered by a random  $x$  such that  $x \in \mathcal{T}_{\{2(i-1)+1\}}$ .
- The condition  $\mathcal{B} = B_0, B_1, \dots$  fails at the first query  $\pi_i(x)$  such that there exists  $y$  which satisfies  $(x, y) \in \mathcal{T}_{\{2(i-1)+1, 2(i-1)+2\}} \setminus \mathcal{P}_i$ , or  $\pi_i^{-1}(y)$  such that there exists  $x$  satisfying  $(x, y) \in \mathcal{T}_{\{2(i-1)+1, 2(i-1)+2\}} \setminus \mathcal{P}_i$ .

By the above representations of  $\mathbf{S}$  and  $\mathbf{T}$ , one can easily verify that  $\mathbf{S}'|\mathcal{A} \equiv \mathbf{S}$  and  $\mathbf{S}'|\mathcal{B} \equiv \mathbf{T}$ . Therefore, by the triangle inequality and by claim (i) from Lemma 1,

$$\Delta_h(\mathbf{S}, \mathbf{T}) \leq \Delta_h(\mathbf{S}, \mathbf{S}') + \Delta_h(\mathbf{S}', \mathbf{T}) \leq \nu(\mathbf{S}', \overline{A}_h) + \nu(\mathbf{S}', \overline{B}_h).$$

To upper bound  $\nu(\mathbf{S}', \overline{A}_h)$ , note that each time a fresh random value is chosen from  $\{0, 1\}^n \setminus \mathcal{P}_{i,j}$  when answering the  $i^{\text{th}}$  query, it is in  $\mathcal{T}_{2(i-1)+j}$  with probability at most  $\frac{i-1}{2^n-i} \leq 2\frac{i-1}{2^n}$ , hence the union bound gives us  $\nu(\mathbf{S}', \overline{A}_h) \leq \frac{h^2}{2^n}$ .

In order to bound  $\nu(\mathbf{S}', \overline{B}_h)$ , let us introduce a monotone condition  $\mathcal{C} = C_0, C_1, \dots$  on  $\mathbf{T}$  which fails under the same circumstances as  $\mathcal{B}$  in  $\mathbf{S}'$  (note that this can be done since  $\mathbf{T}$  also keeps track of the sets  $\mathcal{T}$  and  $\mathcal{P}_i$ ). As a consequence of these equivalent definitions and the fact that the behaviors of  $\mathbf{S}'$  and  $\mathbf{T}$  are the same as long as the respective associated conditions are satisfied, we have  $\nu(\mathbf{S}', \overline{B}_h) = \nu(\mathbf{T}, \overline{C}_h)$ . However, the input-output behavior of  $\mathbf{T}$  is independent of  $Z$  (and  $\mathcal{C}$  failing), and hence we can equivalently postpone the sampling of  $Z$  to the end of the interaction, go through the generated transcript to construct  $\mathcal{T}$ , and upper bound the probability that  $\mathcal{C}$  has

failed at some query. This implies that for the choice of  $Z$ , one query must have been *bad* in the following sense:

- A query  $\pi_1(x)$  is preceded by a  $\pi_2$ -query resulting in an input-output pair  $(x', y')$  such that  $y' \oplus Z = x$ ;
- A query  $\pi_1^{-1}(y)$  is preceded by a  $\pi_2$ -query resulting in an input-output pair  $(x', y')$  such that  $x' \oplus Z = y$ ;
- A query  $\pi_2(x')$  is preceded by a  $\pi_1$ -query resulting in an input-output pair  $(x, y)$  such that  $y \oplus Z = x'$ ;
- A query  $\pi_2^{-1}(y')$  is preceded by a  $\pi_1$ -query resulting in an input-output pair  $(x, y)$  such that  $x \oplus Z = y'$ .

Given the transcript, and for randomly chosen  $Z$ , the  $i^{\text{th}}$  query is bad with probability at most  $(i-1)/2^n$ , and the probability that at least one query is bad is thus at most  $\frac{h^2}{2^{n+1}}$  by the union bound.

Putting all the obtained terms together, the part of the distinguisher's advantage that depends on  $h$  is  $f(h) = q/h2^{\kappa-1} + 3h^2/2^{n+1}$ . This term is minimal for  $h^* = (\frac{1}{3}q2^{n-\kappa+1})^{1/3}$  which gives us  $f(h^*) < 4 \cdot \left(\frac{q}{2^{\kappa+n/2}}\right)^{2/3}$  as desired.  $\square$

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## A Attacks against One-Query Constructions

In this appendix we give the proofs of all claims presented in Section 3.1. To capture the full generality of our argument, we prove the underlying lemmas in a slightly more general scenario. Instead of an ideal block cipher  $\mathbf{E}$ , we consider an arbitrary random function  $\mathbf{F}: \{0, 1\}^d \rightarrow \{0, 1\}^*$  as the available component for the construction. The presented lemmas also hold in this more general setting, which might be of independent interest. Clearly, the setting from Section 3.1 is easily obtained by considering  $d := \kappa + n + 1$  (the input size of the ideal blockcipher  $\mathbf{E}$ ).

Recall that  $\mathbf{C}^{(\cdot)}$  denotes a stateless and deterministic construction invoking the random function  $\mathbf{F}$  to implement a function  $\mathbf{C}^{\mathbf{F}}: \{0, 1\}^{\kappa'} \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  to serve as a block cipher with key length  $\kappa'$  (in particular,  $\mathbf{C}^{\mathbf{F}}$  realizes a permutation for each  $k' \in \{0, 1\}^{\kappa'}$ ). Note that since the construction  $\mathbf{C}^{(\cdot)}$  itself is deterministic and stateless, any adversary with access to  $\mathbf{F}$  can compute the output of  $\mathbf{C}^{\mathbf{F}}$  on any input from  $\{0, 1\}^{\kappa'} \times \{0, 1\}^n \times \{+, -\}$  by simulating  $\mathbf{C}^{(\cdot)}$  on its own and querying  $\mathbf{F}$  when necessary. On the other hand, for a random (secret)  $\kappa'$ -bit string  $K'$ ,  $\mathbf{C}_{K'}^{\mathbf{F}}$  denotes the system which gives access to the permutation  $\mathbf{C}^{\mathbf{F}}(K', \cdot)$  and its inverse (i.e., takes inputs from  $\{0, 1\}^n \times \{+, -\}$ ).

We start by proving the following generalization of Lemma 2.

**Lemma 4.** Let  $\mathbf{C}^{(\cdot)}$  be a one-injective-query construction as above, let  $\mathbf{P}$  be a URP on  $\{0, 1\}^n$ , and let  $K' \in \{0, 1\}^{\kappa'}$  be a random key. Then, for all parameters  $0 < t < 2^{\min\{n, d/2\}-1}$ , there exists a distinguisher  $\mathbf{D}$  such that for all random functions  $\mathbf{F} : \{0, 1\}^d \rightarrow \{0, 1\}^*$ ,

$$\Delta^{\mathbf{D}}((\mathbf{F}, \mathbf{C}_{K'}^{\mathbf{F}}), (\mathbf{F}, \mathbf{P})) \geq 1 - 2/t - 2^{\kappa' - t \cdot (n-1)},$$

where  $\mathbf{D}$  makes  $2t \cdot \max\{2^{d/2}, 2^{d-n}\}$  queries to  $\mathbf{F}$  as well as  $\min\{2^{d/2}, 2^n\}$  forward queries to either of  $\mathbf{C}_{K'}^{\mathbf{F}}$  and  $\mathbf{P}$ .

*Proof.* We start by fixing the parameter  $t$  from the given interval. The distinguisher  $\mathbf{D}$  is described in the top part of Figure 2, where

$$q_{\mathbf{S}} := \min\{2^{d/2}, 2^n\} \quad \text{and} \quad q_{\mathbf{F}} := 2t \cdot \max\{2^{d/2}, 2^{d-n}\},$$

$x_1, \dots, x_{q_{\mathbf{S}}} \in \{0, 1\}^n$  are arbitrarily chosen distinct fixed strings and  $r_1, \dots, r_{q_{\mathbf{F}}}$  are randomly chosen distinct elements of  $\{0, 1\}^d$ , as indicated in Figure 2.

For the analysis, assume that  $\mathbf{S} = \mathbf{C}_{K'}^{\mathbf{F}}$  and the key  $K'$  has been chosen to take the value  $k'$ . Then, the expected size of the set  $\mathcal{S} = \{r_1, \dots, r_{q_{\mathbf{F}}}\} \cap \{q(k', x_1), \dots, q(k', x_{q_{\mathbf{S}}})\}$  is at least  $2t$ , with variance upper-bounded also by  $2t$  by Lemma 6 in Appendix D. Therefore, by Chebyshev's inequality, the probability that  $|\mathcal{S}| < t$  is upper-bounded by  $2/t$ . If  $|\mathcal{S}| \geq t$  then the set  $\mathcal{I}$  is determined on Line 6 and the condition on Line 8 is satisfied and  $\mathbf{D}$  outputs 1. (Note that this condition is verified by simulating the construction  $\mathbf{C}_{k'}$  and querying  $\mathbf{F}$  for the necessary values; at most  $q_{\mathbf{F}}$  queries are needed in total to verify the condition for all  $k'$ , due to the choice of  $\mathcal{I}$ .) On the other hand, if  $\mathbf{S} = \mathbf{P}$ , then for every fixed value  $k'$  the condition on Line 8 is only satisfied with probability  $(2^n - t)!/2^{n!} \leq 2^{-t \cdot (n-1)}$ . Hence, by the union bound, the probability that this condition is satisfied for *any* key  $k'$  is upper-bounded by  $2^{\kappa' - t(n-1)}$ . Putting all of this together, we see that  $\mathbf{D}(\mathbf{F}, \mathbf{C}_{K'}^{\mathbf{F}})$  outputs 1 with probability at least  $1 - 2/t$ , whereas  $\mathbf{D}(\mathbf{F}, \mathbf{P})$  outputs 1 with probability at most  $2^{\kappa' - t(n-1)}$ , which yields the statement of the lemma.  $\square$

Next, we prove the following generalization of Lemma 3 for the scenario with the underlying random function  $\mathbf{F}$ , which we require to have a finite range  $\{0, 1\}^r$ .

**Lemma 5.** Let  $\mathbf{C}^{(\cdot)} : \{0, 1\}^{\kappa'} \times \{0, 1\}^n \times \{+, -\} \rightarrow \{0, 1\}^n$  be a one-query construction, let  $\mathbf{P}$  be a URP on  $\{0, 1\}^n$  and let  $\mathbf{F} : \{0, 1\}^d \rightarrow \{0, 1\}^r$  be a random function. Also, let  $K' \in \{0, 1\}^{\kappa'}$  be a random key, and assume that there exists  $q^*$  such that  $q(K') \leq q^*$  with probability at least  $\frac{1}{2}$ . Then, there exists a distinguisher  $\mathbf{D}$  asking  $2^n$  forward queries such that

$$\Delta^{\mathbf{D}}(\mathbf{C}_{K'}^{\mathbf{F}}, \mathbf{P}) \geq \frac{1}{2} - 2^{\kappa' + r \cdot q^* - \log(2^{n!})}.$$

*Proof.* Let  $\mathcal{K}^*$  be the set of keys  $k'$  for which  $q(k') \leq q^*$ . By our assumption,  $\mathbb{P}[K' \in \mathcal{K}^*] \geq \frac{1}{2}$ . Also, let  $\mathcal{M}$  be the set of mappings which can be implemented by  $\mathbf{C}_{K'}^{\mathbf{F}}$ , given  $K' \in \mathcal{K}^*$ . We have  $|\mathcal{M}| \leq |\mathcal{K}^*| \cdot 2^{r \cdot q^*} \leq 2^{\kappa' + r \cdot q^*}$  since given  $K' \in \mathcal{K}^*$ ,  $\mathbf{C}_{K'}^{\mathbf{F}}$  obtains at most  $r \cdot q^*$  output bits from  $\mathbf{F}$ . Note that  $\mathcal{M}$  only depends on  $\mathbf{C}^{(\cdot)}$  and the description of  $\mathbf{F}$ , and can hence be computed by  $\mathbf{D}$ .

The distinguisher  $\mathbf{D}$  queries the given system for all values in  $\{0, 1\}^n \times \{+\}$  and returns 1 if and only if the mapping obtained is in  $\mathcal{M}$ , and 0 otherwise. Obviously,  $\mathbf{D}$  outputs 1 with probability at least  $\frac{1}{2}$  when interacting with  $\mathbf{C}_{K'}^{\mathbf{F}}$ . However, if  $\mathbf{D}$  interacts with  $\mathbf{P}$ , the probability that it observes a mapping from  $\mathcal{M}$  is upper bounded by  $2^{\kappa' + r \cdot q^*} \cdot 2^{-\log(2^{n!})}$ , which concludes the proof.  $\square$

Finally, by combining the proof techniques of Lemma 4 and Lemma 5, we obtain a proof of Theorem 1.

|   |  |
|---|--|
| <b>Distinguisher D(F, S):</b>   | // $\mathbf{S} \in \{\mathbf{C}_{K'}^{\mathbf{F}}, \mathbf{P}\}$ |
| 1: <b>choose arbitrary distinct</b> $x_1, \dots, x_{q_{\mathbf{S}}} \in \{0, 1\}^n$<br>2: <b>for</b> $i := 1$ to $q_{\mathbf{S}}$ <b>do</b><br>3:   query $\mathbf{S}$ on input $(x_i, +)$ , obtaining value $y_i$<br>4: <b>choose</b> $r_1, \dots, r_{q_{\mathbf{F}}} \xleftarrow{\$} \{0, 1\}^d$<br>5: <b>for all</b> $k' \in \{0, 1\}^{\kappa'}$ <b>do</b><br>6:     let $\mathcal{I} \subseteq \{1, \dots, q_{\mathbf{S}}\}$ be smallest such that <div style="text-align: center; margin: 5px 0;"> <math> \{q(k', x_i) : i \in \mathcal{I}\}  = t</math> and <math>\forall i \in \mathcal{I} : q(k', x_i) \in \{r_1, \dots, r_{q_{\mathbf{F}}}\}</math> </div> 7: <b>if</b> $\mathcal{I}$ exists <b>then</b><br>8: <b>if</b> $\forall i \in \mathcal{I} : y_i = \mathbf{C}^{\mathbf{F}}(k', x_i)$ <b>then</b><br>9: <b>return</b> 1<br>10: <b>return</b> 0   |  |
| <b>Distinguisher D(E, S):</b>   | // $\mathbf{S} \in \{\mathbf{C}_{K'}^{\mathbf{E}}, \mathbf{P}\}$ |
| 1: <b>for all</b> $x \in \{0, 1\}^n$ <b>do</b><br>2:   query $\mathbf{S}$ on input $(x, +)$ , obtaining value $y(x)$<br>3: $r_1, \dots, r_{q_{\mathbf{E},1}} \xleftarrow{\$} \{0, 1\}^{\kappa} \times \{0, 1\}^n \times \{+, -\}$<br>4: $u_1, \dots, u_{q_{\mathbf{E},2}} \xleftarrow{\$} \{0, 1\}^{\kappa} \times \{0, 1\}^n \times \{+, -\}$<br>5: <b>for all</b> $r \in \{r_1, \dots, r_{q_{\mathbf{E},1}}, u_1, \dots, u_{q_{\mathbf{E},2}}\}$ <b>do</b><br>6:     query $\mathbf{E}$ on input $r$ obtaining value $e(r)$<br>7: <b>for all</b> $k' \in \{0, 1\}^{\kappa'}$ <b>do</b><br>8:     let $\mathcal{I} \subseteq \{0, 1\}^n$ be smallest such that $ \mathcal{I}  = t$ and $\forall x \in \mathcal{I}$ : <div style="text-align: center; margin: 5px 0;"> <math>q_1(k', x) \in \{r_1, \dots, r_{q_{\mathbf{E},1}}\}</math> and <math>q_2(k', x, e(q_1(k', x))) \in \{u_1, \dots, u_{q_{\mathbf{E},2}}\}</math> </div> 9: <b>if</b> $\mathcal{I}$ exists <b>then</b><br>10: <b>if</b> $\forall x \in \mathcal{I} : y(x) = \mathbf{C}^{\mathbf{E}}(k', x)$ <b>then</b><br>11: <b>return</b> 1<br>12: <b>return</b> 0 |  |

**Fig. 2.** Top: Distinguisher  $\mathbf{D}$  for the proofs of Lemma 4 and Theorem 1. Bottom: Distinguisher  $\mathbf{D}$  for the proof of Theorem 2. In both descriptions, “smallest” in the choice of the set  $\mathcal{I}$  is with respect to the lexicographic order on sets of cardinality  $t$ .

*Proof (of Theorem 1).* Throughout the proof, fix  $q^* = ((n - 2) \cdot 2^n - \kappa' - 1) / n$ . Assume first that  $q(K') \leq q^*$  with probability at least  $\frac{1}{2}$ . Then we can use Lemma 3 to obtain a distinguisher  $\mathbf{D}$  asking  $2^n$  queries such that

$$\Delta^{\mathbf{D}}(\mathbf{C}_{K'}^{\mathbf{E}}, \mathbf{P}) \geq \frac{1}{2} - 2^{\kappa' + n \cdot q^* - \log(2^{n!})} \geq \frac{1}{2} - 2^{(n-2) \cdot 2^n - \log(2^{n!}) - 1}.$$

Applying the bound  $\ln x! \geq x \ln(x/e) + 1$  we get  $\log(2^{n!}) \geq (n - 2) \cdot 2^n + 1$  and therefore  $\Delta^{\mathbf{D}}(\mathbf{C}_{K'}^{\mathbf{E}}, \mathbf{P}) \geq 1/4$ . This also trivially implies that there exists a distinguisher  $\mathbf{D}$  asking  $2^n$  queries such that we have  $\Delta^{\mathbf{D}}((\mathbf{E}, \mathbf{C}_{K'}^{\mathbf{E}}), (\mathbf{E}, \mathbf{P})) \geq 1/4$ .

Let us address the complementary case that  $q(K') > q^*$  with probability at least  $\frac{1}{2}$ . Since we assume  $\kappa' \leq 2^n - 1$  and  $n \geq 6$ , we note that  $q^* \geq 2^n \cdot (n - 3) / n \geq 2^{n-1}$ . Hence, roughly speaking,

for at least half of the keys  $k'$  the mapping  $q(k', \cdot)$  is “almost injective” and we can use the same distinguisher as in the proof of Lemma 4 with different parameters.

More precisely, we use the distinguisher  $\mathbf{D}$  depicted in the top part of Figure 2 with  $q_{\mathbf{S}} := 2^n$  (we query the entire domain) and  $q_{\mathbf{F}} := 8t \cdot 2^\kappa$  (recall that the ideal cipher  $\mathbf{E}$  has domain size  $\kappa + n + 1$ ). If  $k'$  is chosen such that  $q(k') > 2^{n-1}$  (which is true with probability at least  $\frac{1}{2}$ ), then the expected value and the variance of the cardinality of  $\mathcal{S} := \{q(k', x) \mid x \in \{0, 1\}^n \cap \{r_1, \dots, r_{q_{\mathbf{F}}}\}\}$  are at least  $2t$  (again by Lemma 6) and hence Chebyshev’s inequality yields that  $|\mathcal{S}| \geq t$  except with probability at most  $2/t$ . Therefore  $\mathbf{D}(\mathbf{E}, \mathbf{C}_{K'}^{\mathbf{E}})$  outputs 1 with probability at least  $\frac{1}{2} - 2/t$ . However, as in the proof of Lemma 4,  $\mathbf{D}(\mathbf{E}, \mathbf{P})$  outputs 1 with probability at most  $2^{\kappa' - t \cdot (n-1)}$ . This concludes the proof.  $\square$

## B Attacks against Two-Query Constructions

Here we present the details of a generic attack on the broad class of two-query block-cipher constructions described in Theorem 2.

*Proof (of Theorem 2).* The distinguisher  $\mathbf{D}$  is described in the bottom part of Figure 2, with the parameters  $q_{\mathbf{E},1} := 8 \cdot 2^{\kappa+n/2}$  and  $q_{\mathbf{E},2} := 2t \cdot 2^{\kappa+n/2}$ . Let us analyze the interaction of this distinguisher with  $(\mathbf{E}, \mathbf{C}_{K'}^{\mathbf{E}})$ . For every fixed key  $k'$ , the size of the set  $\{q_1(k', x) \mid x \in \{0, 1\}^n\}$  is  $2^n$  due to the injectivity of the mapping  $q_1(k', \cdot)$ . Hence the expected size of the set  $\mathcal{S} := \{q_1(k', x) \mid x \in \{0, 1\}^n \cap \{r_1, \dots, r_{q_{\mathbf{E},1}}\}\}$  is  $4 \cdot 2^{n/2}$  with the variance upper-bounded by  $4 \cdot 2^{n/2}$  by Lemma 6. Chebyshev’s inequality then yields  $\mathbb{P}[|\mathcal{S}| < 3 \cdot 2^{n/2}] \leq 2^{2-n/2}$ .

Let us assume that  $|\mathcal{S}| \geq 3 \cdot 2^{n/2}$ : We need to show that the image  $\mathbf{E}(\mathcal{S})$  of  $\mathcal{S}$  under  $\mathbf{E}$  is also sufficiently large. Indeed, note that the set  $\mathcal{S}$  is independent of  $\mathbf{E}$ : If an ideal block cipher is queried at  $3 \cdot 2^{n/2}$  arbitrary distinct triples  $(k, x, \sigma) \in \{0, 1\}^\kappa \times \{0, 1\}^n \times \{-, +\}$ , then for any two such triples  $(k, x, \sigma)$  and  $(k', x', \sigma')$ , the probability that the query outputs collide under  $\mathbf{E}$  is at most  $\frac{2}{2^n}$ . (The case with the highest probability is when  $k = k'$ ,  $x \neq x'$ ,  $\sigma = -$ , and  $\sigma' = +$ .) The expected number of collisions is upper-bounded by  $\frac{9 \cdot 2^n}{2} \cdot \frac{2}{2^n} = 9$ . Therefore, by Markov’s inequality no more than  $2^{n/2}$  collisions (among the values  $\mathbf{E}(s)$  for  $s \in \mathcal{S}$ ) occur, except with probability at most  $9 \cdot 2^{-n/2}$ , and if at most  $2^{n/2}$  collisions occur, then  $|\mathbf{E}(\mathcal{S})| \geq 3 \cdot 2^{n/2} - 2^{n/2} = 2^{n/2+1}$ . Due to the assumed property of  $q_2$ , this also means that the set  $\mathcal{T} := \{q_2(k', x, e(q_1(k', x))) : q_1(k', x) \in \mathcal{S}\}$  has cardinality  $|\mathcal{T}| \geq 2^{n/2+1}$ .

Similar to above, the estimated size of the set  $\mathcal{V} := \mathcal{T} \cap \{u_1, \dots, u_{q_{\mathbf{E},2}}\}$  is at least  $2^{n/2+1} \cdot 2t \cdot 2^{\kappa+n/2} / 2^{\kappa+n+1} = 2t$  with variance at most  $2t$  (again by Lemma 6) and hence the probability that  $|\mathcal{V}| < t$  can again be upper-bounded by  $2/t$  using the Chebyshev inequality.

Note that if  $|\mathcal{V}| \geq t$  then a set  $\mathcal{I}$  will be found on Line 8 when going through the right chosen key  $k'$ , which implies that the test on Line 10 succeeds, and we are hence guaranteed that  $\mathbf{D}$  outputs 1, unless one of the previously mentioned bad events happens, for which the probability is bounded by  $13 \cdot 2^{-n/2} + 2/t$ . On the other hand, if  $\mathbf{D}$  interacts with  $\mathbf{P}$ , the probability that this test is satisfied for any of the keys  $k' \in \{0, 1\}^{\kappa'}$  can be upper-bounded by  $2^{\kappa' - t \cdot (n-1)}$  as in the proof of Lemma 4.  $\square$

## C Attacks against Randomized Cascades

To justify the minimality of our approach, we briefly discuss some related constructions and their security: Namely, we show that for various simple key-dependent randomizations of the cascade, meet-in-the-middle attacks issuing only  $2^\kappa$  queries can be mounted.

SINGLE RANDOMIZATION. We start with constructions with one single randomization step: These are  $\mathbf{C}_1$ ,  $\mathbf{C}_2$ , and  $\mathbf{C}_3$  such that

$$\begin{aligned}\mathbf{C}_1^E_{k_1, k_2, z}(m) &= E_{k_2}(E_{k_1}(m \oplus z)) , \\ \mathbf{C}_2^E_{k_1, k_2, z}(m) &= E_{k_2}(E_{k_1}(m)) \oplus z , \\ \mathbf{C}_3^E_{k_1, k_2, z}(m) &= E_{k_2}(E_{k_1}(m) \oplus z) .\end{aligned}$$

For example, given an ideal block cipher  $\mathbf{E}$ , we distinguish  $(\mathbf{E}, \mathbf{C}_3^{\mathbf{E}}_{K_1, K_2, Z})$  (for random  $K_1, K_2$  and  $Z$ ) and  $(\mathbf{E}, \mathbf{P})$  with advantage  $\frac{1}{2}$  using  $\mathcal{O}(2^\kappa)$  queries as follows: Let  $x_1, x_2, \dots, x_t$  be fixed distinct  $n$ -bit strings where  $t = \lceil \frac{2\kappa+n}{n-1} \rceil$ . For all  $i = 1, \dots, t$  and  $k'_1, k'_2 \in \{0, 1\}^\kappa$ , we first compute  $u_i(k'_1)$  and  $v_i(k'_2)$  by querying  $(k'_1, x_i, +)$  and  $(k'_2, \pi(x_i), -)$  to  $\mathbf{E}$ , where  $\pi$  is the permutation implemented by the system  $\mathbf{S} \in \{\mathbf{C}_3^{\mathbf{E}}_{K_1, K_2, Z}, \mathbf{P}\}$  appearing as second component. Finally, we output 1 if there exist  $z' \in \{0, 1\}^n$  and  $k'_1, k'_2 \in \{0, 1\}^\kappa$  with  $u_i(k'_1) \oplus v_i(k'_2) = z'$  for all  $i = 1, \dots, t$ , and 0 otherwise.

Clearly, we will always output one when interacting with  $\mathbf{S} = \mathbf{C}_3^{\mathbf{E}}_{K_1, K_2, Z}$ . However, if  $\mathbf{S} = \mathbf{P}$ , then for each  $k'_1, k'_2$  and  $z$ , the sequence  $v_i(k'_2)$  is a uniformly random  $t$ -tuple of distinct  $n$ -bit strings. Hence, the probability that  $u_i(k'_1) \oplus v_i(k'_2) = z$  for all  $i = 1, \dots, t$  is at most  $(2^n - t)!/2^{nt} < 2^{-t(n-1)}$ , and the probability that there exist such  $k'_1, k'_2$  and  $z$  is hence at most  $2^{2\kappa+n-t(n-1)} \leq \frac{1}{2}$  by the union bound.

Similarly, for  $\mathbf{C}_1$  one modifies the attack to compute  $u_i(k'_1)$  as the result of querying  $\mathbf{E}$  with input  $(k'_1, x_i, -)$ , whereas  $v_i(k'_2)$  is obtained by applying  $\pi^{-1}$  to the output of the query  $(k'_2, x_i, +)$ . To attack  $\mathbf{C}_2$ , we proceed symmetrically.

DOUBLE RANDOMIZATION. It is slightly more complicated to attack the construction

$$\mathbf{C}_4^E_{k_1, k_2, z_1, z_2}(m) = E_{k_2}(E_{k_1}(m \oplus z_1)) \oplus z_2 .$$

We can only show how to do it with  $\mathcal{O}(2^\kappa)$  ideal block cipher queries and  $2^n$  construction queries. Following on the above notation, we let  $u_i(k'_1)$  and  $v_i(k'_2)$  be the results of querying  $\mathbf{E}$  with  $(k'_1, x_i, -)$  and  $(k'_2, x_i, +)$  for all  $k'_1, k'_2 \in \{0, 1\}^\kappa$  and  $i = 1, \dots, m$ , but also query all of the given permutation  $\pi$  (this costs us  $2^n$  queries). Then, as above we check for the existence of  $k'_1, k'_2 \in \{0, 1\}^\kappa$  and  $z_1, z_2 \in \{0, 1\}^n$  such that  $\pi(z_1 \oplus u_i(k'_1)) = v_i(k'_2) \oplus z_2$  for all  $i = 1, \dots, t$ . Note that the probability that this is achieved when  $\pi$  is chosen randomly is at most  $2^{2\kappa+2n} \cdot \frac{(2^n-t)!}{2^{nt}} \leq \frac{1}{2}$  if we set  $t = \lceil \frac{2\kappa+2n}{n-1} \rceil$ . A similar attack can be given when  $z_1 = z_2$ .

## D Estimating Intersection Sizes

The situation analyzed in the following lemma occurs at several occasions during our proofs. Let  $\mathbf{E}$ ,  $\mathbf{Var}$  and  $\mathbf{Cov}$  denote the usual notions of expected value, variance and covariance, respectively.

**Lemma 6.** *Let  $a, b \leq N$ , let  $\mathcal{U}$  be a set such that  $|\mathcal{U}| = N$ , and let  $\mathcal{B}$  be its subset such that  $|\mathcal{B}| = b$ . Now let  $\mathcal{A}$  be chosen uniformly at random from all subsets of  $\mathcal{U}$  of size  $a$ . If the random variable  $X$  denotes the number of elements of the intersection  $\mathcal{A} \cap \mathcal{B}$ , then we have  $\mathbf{E}(X) = ab/N$  and  $\mathbf{Var}(X) \leq ab/N$ .*

*Proof.* First, let us derive the expected value of  $X$ . We can see the set  $\mathcal{A}$  as determined in the following way: we choose a mapping  $F$  uniformly at random from the set of all injective mappings from  $\{1, \dots, a\}$  to  $\mathcal{U}$  and then let  $\mathcal{A}$  be the range of  $F$ . Now we have  $X = \sum_{i=1}^a X_i$  where  $X_i$  is the indicator random variable set to 1 iff  $F(i) \in \mathcal{B}$ . It is easy to see that  $\mathbf{E}(X_i) = b/N$  and by the linearity of expectation we have  $\mathbf{E}(X) = \sum_{i=1}^a \mathbf{E}(X_i) = ab/N$ .

For the variance, let us consider the indicator random variables  $X_i$  again. We have

$$\text{Var}(X) = \sum_{i=1}^a \text{Var}(X_i) + 2 \cdot \sum_{1 \leq i < j \leq a} \text{Cov}(X_i, X_j)$$

and it remains to estimate the terms in this equation. We have

$$\begin{aligned} \text{Var}(X_i) &= \mathbb{E}(X_i^2) - (\mathbb{E}(X_i))^2 = \frac{b}{N} - \left(\frac{b}{N}\right)^2 = \frac{b(N-b)}{N^2} \leq \frac{b}{N} \\ \text{Cov}(X_i, X_j) &= \mathbb{E}(X_i \cdot X_j) - \mathbb{E}(X_i) \cdot \mathbb{E}(X_j) = \frac{b(b-1)}{N(N-1)} - \left(\frac{b}{N}\right)^2 < 0 \end{aligned}$$

hence we obtain the desired result. □